

# Concentration Inequalities and Graph Coloring

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# A Simple Start

Thm (Markov's Inequality)

For  $t > 0$ , we have

$$P(|X| \geq t) \leq \frac{E|X|}{t}$$

pf:

$$\begin{aligned} E|X| &= \int_{\Omega} |X| d\mathcal{P} \geq \int_{|X| \geq t} |X| d\mathcal{P} \\ &\geq t \mathcal{P}(|X| \geq t) \end{aligned}$$

QED



# Chernoff's Inequality

Thm:  $X_1, X_2, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} B(1, p)$ .

Let  $\lambda = np$  and  $S_n = \sum_{i=1}^n X_i$ .

Then, for  $0 \leq t \leq n - \lambda$ ,

$$P\{S_n \geq \lambda + t\} \leq \left(\frac{\lambda}{\lambda + t}\right)^{\lambda + t} \left(\frac{n - \lambda}{n - \lambda - t}\right)^{n - \lambda - t}.$$



pf: For  $u \geq 0$

$$P_r \{ S_n \geq E S_n + t \} \leq e^{-u(\lambda+t)} \prod_{i=1}^n E e^{uX_i}$$

$$= e^{-u(n+t)} (1-p+pe^u)^n \quad *$$

$$\text{let } f(x) = x^{-(\lambda+t)} (1-p+px)^n$$

$$f'(x) = -(\lambda+t) x^{-(\lambda+t+1)} (1-p+px)^n + np x^{-(\lambda+t)} (1-p+px)^{n-1}$$

$$\text{and } f'(x) = 0 \Rightarrow \hat{x} = \frac{(1-p)(\lambda+t)}{(n-\lambda-t)p}$$

$f(x)$  attains its minimum at  $\hat{x}$ . assume  $n-\lambda-t > 0$   
ie  $*$  attains its minimum at  $e^u = \hat{x}$  "

This yields

$$\begin{aligned} P\{S_n \geq ES_{n+t}\} &\leq \left( \frac{(\lambda+t)(1-p)}{(n-\lambda-t)p} \right)^{-(\lambda+t)} \left( 1-p + \frac{(1-p)(\lambda+t)}{(n-\lambda-t)} \right)^n \\ &= \left( \frac{\lambda+t}{n-\lambda-t} \right)^{-(\lambda+t)} \left( \frac{1-p}{p} \right)^{-(\lambda+t)} (1-p)^n \left( \frac{n}{n-\lambda-t} \right)^n \end{aligned}$$

Next we use the fact i.e.  $p = \frac{\lambda}{n}$  to get

$$P\{S_n \geq ES_{n+t}\} \leq \left( \frac{\lambda}{\lambda+t} \right)^{\lambda+t} \left( \frac{n-\lambda}{n-\lambda-t} \right)^{n-\lambda-t} \quad \text{as } 0 \leq t \leq n-\lambda.$$

QED

# Chernoff's bounding technique

If  $\epsilon$  is an arbitrary positive number  
then for any r.v.  $X$  and any  $t > 0$ ,

$$P(X \geq t) = P(e^{\epsilon X} \geq e^{\epsilon t}) \leq \frac{E e^{\epsilon X}}{e^{\epsilon t}}.$$

In Chernoff's method, we find  $\epsilon > 0$  that  
makes the upper bound small.



Extensions of Chernoff's Ineq. can be derived from....

Lemma\*: Let  $X_1, \dots, X_n$  be independent with  $0 \leq X_i \leq 1$  for each  $i$ . Let  $p = \frac{ES_n}{n}$ , where  $S_n = \sum_{i=1}^n X_i$ .

Then for any  $0 \leq \tau < 1-p$ ,

$$P(S_n - ES_n \geq n\tau) \leq \left( \left( \frac{p}{p+\tau} \right)^{p+\tau} \left( \frac{1-p}{1-p-\tau} \right)^{1-p-\tau} \right)^n$$

pf:  $P(S_n \geq np + nt) \leq e^{-u(np+nt)} \prod_{i=1}^n \mathbb{E} e^{uX_i}$ , where  $u > 0$

$$\leq e^{-u(np+nt)} \prod_{i=1}^n \mathbb{E} (1 - X_i + X_i e^u)$$

$$\leq e^{-u(np+nt)} \left( \frac{n - \mathbb{E} S_n + (\mathbb{E} S_n) e^u}{n} \right)^n$$

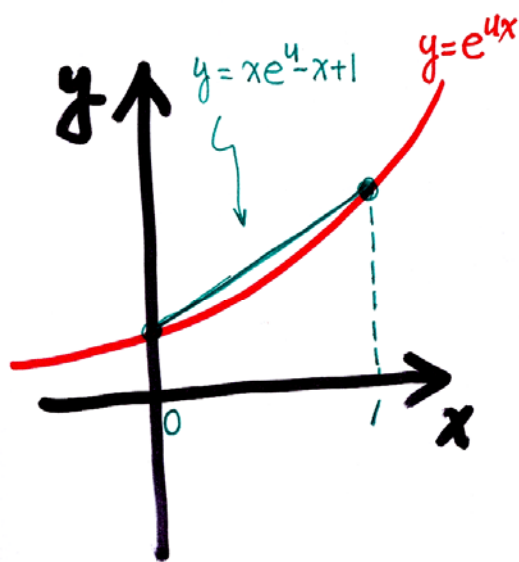
geometric mean  $\leq$  arithmetic mean

$$= \left[ e^{-u(p+t)} (1 - p + p e^u) \right]^n$$

= RHS by letting

$$e^u = \frac{(p+t)(1-p)}{p(1-p-t)}$$

**QED**





## Weaker but more useful bounds

Thm: Let  $X_1, X_2, \dots, X_n$  be independent rvs s.t.  
 $0 \leq X_i \leq 1$  for each  $i$ . Let  $p = \frac{ES_n}{n}$ , where  
 $S_n = \sum_{i=1}^n X_i$ . Then

(a) For any  $\epsilon > 0$ ,

$$\begin{aligned} P(S_n \geq (1+\epsilon)np) &\leq \exp(-np((1+\epsilon)\ln(1+\epsilon) - \epsilon)) \\ &\leq \exp\left(-\frac{\epsilon^2 np}{2(1+\epsilon/3)}\right) \end{aligned}$$

(b) For any  $\epsilon > 0$ ,

$$P(S_n \leq (1-\epsilon)np) \leq \exp\left(-\frac{1}{2}\epsilon^2 np\right)$$



pf of (a) In the proof of Lemma\*,

let  $\delta = \epsilon p$  and  $e^u = (1 + \epsilon)$ , then we have

$$P(S_n \geq (1 + \epsilon)np)$$

$$= P(S_n \geq np + n\delta)$$

$$\leq \left[ \frac{(1 + \epsilon)^{-(1 + \epsilon)}}{(1 + \epsilon p)} \right]^n \quad (\text{by the proof of Lemma*})$$

$$\leq \left[ \frac{(1 + \epsilon)^{-(1 + \epsilon)}}{(1 + \epsilon p)^{\frac{1}{p}}} \right]^{np}$$

$$\leq \left[ (1 + \epsilon)^{-(1 + \epsilon)} e^{\epsilon} \right]^{np} \quad (\because 1 + \epsilon p \leq e^{\epsilon p})$$

this proves the first inequality in (a).

pf of (a): Claim For all  $x \geq 0$ ,  $(1+x) \ln(1+x) - x \geq \frac{3x^2}{6+2x}$ .

Note that  $\star = e^{-np[(1+\epsilon) \ln(1+\epsilon) - \epsilon]}$   
 $\leq e^{-np \frac{3\epsilon^2}{6+2\epsilon}}$ , done.

pf of (b): Let  $f(t) = \ln \left( \left( \frac{p}{p+t} \right)^{p+t} \left( \frac{1-p}{1-p-t} \right)^{1-p-t} \right)$

Let  $h(x) = f(-xp)$  for  $0 \leq x < 1$ .

Then  $h(0) = f(0) = 0$ ,  $h'(0) = f'(0)(-p) = 0$  ( $\because f'(t) = \ln \left( \frac{p(1-p-t)}{(p+t)(1-p)} \right)$ )

$$h''(x) = f''(-xp)p^2 = -\frac{p}{(1-x)(1-p+xp)} \leq -p \quad (\because 0 \leq 1-x \leq 1, 0 \leq 1-p(1-x) \leq 1)$$

(Note  $f''(t) = -\frac{1}{(p+t)(1-p-t)}$ )

$$\begin{aligned} \text{Taylor's Thm} \Rightarrow h(x) &= h(0) + h'(0)x + \frac{h''(\theta)x^2}{2}, \quad 0 \leq \theta \leq x \\ &\leq 0 + 0 - \frac{px^2}{2} \end{aligned}$$



pf of (b) Lemma\* implies that

$$P((n-S_n) - (n-\varepsilon S_n) \geq nt) \leq e^{f(t)n}$$

let  $t = \varepsilon p$ . we have

$$\begin{aligned} P(S_n \leq (1-\varepsilon)\varepsilon S_n) &= P(S_n \leq \varepsilon S_n - nt) \leq e^{f(t)n} \\ &= e^{h(-\frac{t}{p})n} \leq e^{-\frac{p}{2}(\frac{t}{p})^2 n} = e^{-\frac{1}{2}\varepsilon^2 np} \end{aligned}$$

$$(\because h(x) \leq -\frac{px^2}{2}, \text{ for } 0 \leq x < 1)$$

**QED**



Remark: The first inequality in (a) implies

$$P(S_n \geq 2\epsilon S_n) \leq e^{-(0.386)\epsilon S_n} \quad \text{and}$$

$$P(S_n \geq \delta \epsilon S_n) \leq e^{-\delta \ln(\frac{\delta}{e})\epsilon S_n}$$

The second inequality in (a) implies

$$P(S_n \geq (1+\epsilon)\epsilon S_n) \leq e^{-\frac{1}{3}\epsilon^2 \epsilon S_n}$$

# Hoeffding's Lemma

Lemma:  $E X = 0$ ,  $a \leq X \leq b$ .

Then for any  $\lambda > 0$ ,

$$E(e^{\lambda X}) \leq e^{\frac{\lambda^2(b-a)^2}{8}}$$



**pf:** Note that  $e^{px} \leq \left(\frac{x-a}{b-a}\right)(e^{pb} - e^{pa}) + e^{pa}$

for  $a \leq x \leq b$ . It follows that

$$xe^{px} \leq \frac{-a}{b-a}(e^{pb} - e^{pa}) + e^{pa}$$

$$= e^{\phi(u)}, \text{ where } u \stackrel{\text{def}}{=} p(b-a), \phi(u) \stackrel{\text{def}}{=} -pu + \ln(1-p+pe^u)$$

$$\text{and } p \stackrel{\text{def}}{=} -\frac{a}{b-a}$$



$$\text{Note that } \phi'(u) = -p + \frac{p}{p+(1-p)e^{-u}}, \phi''(u) = \frac{p(1-p)e^{-u}}{(p+(1-p)e^{-u})^2} \leq \frac{p(1-p)\frac{p}{1-p}}{(p+(1-p)\frac{p}{1-p})^2} = \frac{1}{4}$$

Taylor's Thm  $\Rightarrow$  for some  $\theta \in [0, u]$ ,

$$\phi(u) = \underbrace{\phi(0)}_{=0} + \underbrace{\phi'(0)}_{=0}u + \frac{\phi''(\theta)u^2}{2} \leq \frac{u^2}{8} = \frac{p^2(b-a)^2}{8}$$

**QED**



# Hoeffding's Inequality

Thm

Let  $X_1, X_2, \dots, X_n$  be independent rvs s.t.

$a_i \leq X_i \leq b_i$  with probability one,  $1 \leq i \leq n$ .

Then for any  $t \geq 0$  we have

$$\mathcal{P}(S_n - \mathbb{E} S_n \geq t) \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right)$$

$$\mathcal{P}(S_n - \mathbb{E} S_n \leq -t) \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right)$$

Where  $S_n = \sum_{i=1}^n X_i$ .

$$\text{pt: } P(S_n - ES_n \geq t) = P(e^{\lambda(S_n - ES_n)} \geq e^{\lambda t})$$

$$\leq e^{-\lambda t} E e^{\lambda(S_n - ES_n)}$$

$$= e^{-\lambda t} \prod_{i=1}^n E e^{\lambda(X_i - EX_i)}$$

$$\leq e^{-\lambda t} \prod_{i=1}^n e^{-\frac{\lambda^2 (b_i - a_i)^2}{8}} \quad \text{by Hoeffding's Lemma}$$

$$= e^{-\lambda t} e^{-\frac{\lambda^2}{8} \sum_{i=1}^n (b_i - a_i)^2} \quad \begin{aligned} & \because (b_i - EX_i) - (a_i - EX_i) \\ & = b_i - a_i \end{aligned}$$

$$= e^{-\frac{2\lambda^2 t^2}{\sum_{i=1}^n (b_i - a_i)^2}} \quad \text{(by choosing } \lambda = \frac{4t}{\sum_{i=1}^n (b_i - a_i)^2} \text{)}$$



# Chebyshev's Ineq. is Inadequate

Let  $X_1, X_2, \dots \stackrel{iid}{\sim} B(1, p)$

Chebyshev's Inequality says that  $P(|S_n - ES_n| \geq n\epsilon) \leq \frac{p(1-p)}{n\epsilon^2}$

However, as  $n \rightarrow \infty$ , Central Limit Thm suggests that

$$P(S_n - ES_n \geq n\epsilon) = P\left(\frac{S_n - ES_n}{\sqrt{\text{Var} S_n}} \geq \frac{n\epsilon}{\sqrt{np(1-p)}}\right) \sim 1 - \Phi\left(\frac{n\epsilon}{\sqrt{np(1-p)}}\right) \\ \leq \exp\left(-\frac{n\epsilon^2}{2p(1-p)}\right) \quad \left(\because \sqrt{2\pi} \{1 - \Phi(x)\} < \frac{1}{x} \exp\left(-\frac{x^2}{2}\right)\right)$$

It seems Hoeffding Ineq. is adequate .....

# However....

Hoeffding's bound is independent of  $p$ , so it may loose if  $p$  is small (or large).

That is this inequality ignores information about the variance of the  $X_i$ 's.

The Bernstein's Inequality will give an answer.....



# A Lemma first

Lemma: If  $|X| \leq c$ ,  $\mathbb{E}X = 0$  and  $\text{Var}(X) = \sigma^2$ ,

then 
$$\mathbb{E}(e^{sX}) \leq \exp\left(\frac{\sigma^2}{c}(e^{sc} - 1 - sc)\right)$$

Pf:  $e^{Ax} = 1 + Ax + \sum_{r=2}^{\infty} \frac{(Ax)^r}{r!}$

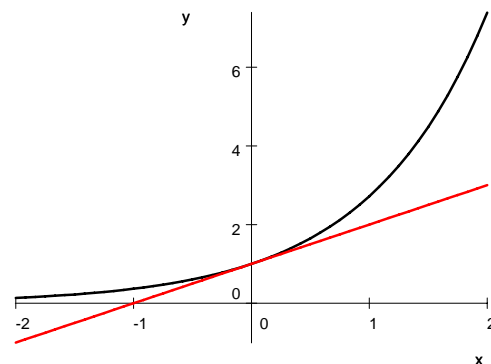
$$\Rightarrow E e^{AX} = 1 + \sum_{r=2}^{\infty} \frac{A^r EX^r}{r!} \quad (\because EX=0)$$

$$\leq 1 + \sum_{r=2}^{\infty} \frac{A^r C^{r-2} \sigma^2}{r!} \quad \left( \because EX^r \leq E|X|^{r-2} X^2 \leq C^{r-2} \sigma^2 \text{ for } r \geq 2 \right)$$

$$= 1 + \frac{\sigma^2}{C^2} \sum_{r=2}^{\infty} \frac{(AC)^r}{r!}$$

$$= 1 + \frac{\sigma^2}{C^2} (e^{AC} - 1 - AC)$$

$$\leq e^{\frac{\sigma^2}{C^2} (e^{AC} - 1 - AC)} \quad (\because e^x \geq 1+x \quad \forall x)$$





# The Bennett's Inequality

Thm Let  $X_1, X_2, \dots, X_n$  be independent rvs with  $E X_i = 0$  and  $|X_i| \leq c$ ,  $1 \leq i \leq n$ . Let  $S_n = \sum_{i=1}^n X_i$  and  $\sigma^2 = \frac{1}{n} \sum_{i=1}^n \text{Var}(X_i)$ . Then for any  $t > 0$ ,

$$P(S_n > t) \leq \exp\left(-\frac{t}{c} \left( \left(1 + \frac{n\sigma^2}{ct}\right) \ln\left(1 + \frac{ct}{n\sigma^2}\right) - 1 \right)\right)$$

# The Proof

Proof:

$$\begin{aligned} P(S_n > t) &\leq e^{-\lambda t} E e^{\lambda S_n} && \text{(Markov's inequality)} \\ &= e^{-\lambda t} \prod_{i=1}^n E e^{\lambda X_i} \\ &\leq e^{-\lambda t} \prod_{i=1}^n e^{\frac{\sigma_i^2}{c^2} (\lambda c - \lambda^2 c)} && \text{by above lemma} \\ &= e^{-\lambda t} e^{\frac{n \sigma^2}{c^2} (\lambda c - \lambda^2 c)} \end{aligned}$$

The last term is minimized for

$$\lambda = \frac{1}{c} \ln \left( 1 + \frac{tc}{n\sigma^2} \right).$$

Resubstituting this value, we obtain Bennett's inequality.



# Bernstein's Inequality

Thm Under the conditions of Bennett's Ineq, for  $\varepsilon > 0$ , we have

$$P(S_n > n\varepsilon) \leq \exp\left(-\frac{n\varepsilon^2}{2\sigma^2 + c\varepsilon}\right)$$

pf: Applying the following elementary ineq.

to Bennett's bound:  $\ln(1+x) \geq \frac{2x}{2+x}, x \geq 0$

Bennett's bound

$$\ln(1+x) = x - (1/2)x^2 + (1/3)x^3 + O(x^4)$$

$$(2x)/(2+x) = x - (1/2)x^2 + (1/4)x^3 + O(x^4)$$

$$\leq \exp \left( -\frac{n\varepsilon}{c} \left( \left( 1 + \frac{n\sigma^2}{c\varepsilon n} \right) \left( \frac{2 \frac{c n \varepsilon}{n \sigma^2}}{2 + \frac{c n \varepsilon}{n \sigma^2}} \right) - 1 \right) \right)$$

$$= \exp \left( -\frac{n\varepsilon}{c} \left( \frac{c\varepsilon + \sigma^2}{c\varepsilon} \frac{2c\varepsilon}{c\varepsilon + 2\sigma^2} - 1 \right) \right)$$

$$= \exp \left( -\frac{n\varepsilon}{c} \frac{c\varepsilon}{2\sigma^2 + c\varepsilon} \right) = \exp \left( -\frac{n\varepsilon^2}{2\sigma^2 + c\varepsilon} \right)$$

QED



What's after Bernstein's ineq.....

## Poisson-type Inequality

Thm Let  $X_1, X_2, \dots, X_n$  be independent rvs  
with  $0 \leq X_i \leq 1 \quad \forall i$  and  $m = \mathbb{E}S_n$ .

Then for any  $t \geq m$ ,

$$P(S_n \geq t) \leq \left(\frac{m}{t}\right)^t e^{t-m}$$

Note: Here  $n$  doesn't appear on the RHS

pf: let  $m_i = EX_i$ .

$$P(S_n \geq t) \leq e^{-st} \prod_{i=1}^n E e^{sX_i}$$

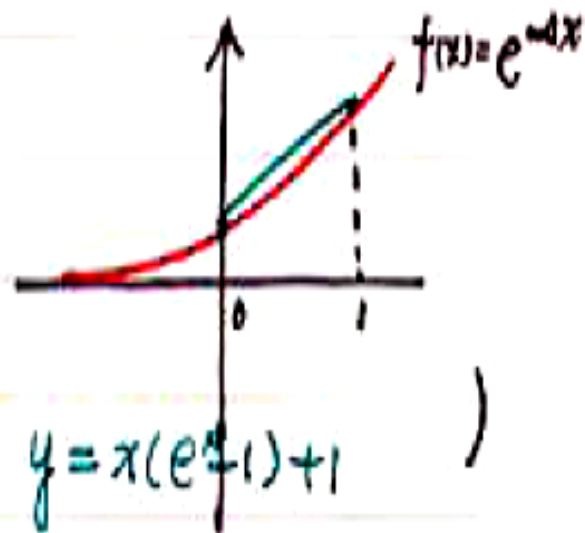
$$\leq e^{-st} \prod_{i=1}^n E(X(e^s - 1) + 1) \quad (\because$$

$$= e^{-st} \prod_{i=1}^n [1 + m_i(e^s - 1)]$$

$$\leq e^{-st} \prod_{i=1}^n e^{m_i(e^s - 1)} = e^{-st} e^{m(e^s - 1)} \star$$

$$\leq \left(\frac{m}{t}\right)^t e^{t-m} \text{ by choosing } s = \ln\left(\frac{t}{m}\right)$$

to minimize the bound  $\star$  QED



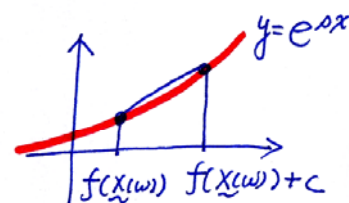


# An extension of Hoeffding's Lemma

Lemma: Suppose for r.v.  $V$  and r.vector  $\underline{X}$  we have

①  $E(V | \underline{X}) = 0$  a.s. ②  $f(\underline{x}) \leq V \leq f(\underline{x}) + c$ , for some fun.  $f$  and constant  $c$ .

Then for any  $\lambda > 0$ ,  $E(e^{\lambda V} | \underline{X}) \leq e^{\frac{\lambda^2 c^2}{8}}$ .



pf: Note that

$$e^{\lambda V} \leq \frac{V - f(\underline{x})}{c} (e^{\lambda(f(\underline{x}) + c)} - e^{\lambda f(\underline{x})}) + e^{\lambda f(\underline{x})}$$

$$\text{Thus } E(e^{\lambda V} | \underline{X}) = \frac{V - f(\underline{x})}{c} (e^{\lambda(f(\underline{x}) + c)} - e^{\lambda f(\underline{x})}) + e^{\lambda f(\underline{x})}$$

The remains is similar to Hoeffding's Lemma.

# McDiarmid's Inequality

Thm: Let  $X_1, X_2, \dots, X_n$  be independent rvs. Let  $f$  be a function  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  with a vector  $(c_1, c_2, \dots, c_n)$  s.t.  $|f(\underline{x}) - f(\underline{y})| \leq c_i$  for all  $\underline{x}, \underline{y}$  in  $\mathbb{R}^n$  that differ only at the  $i$ th coordinate,  $1 \leq i \leq n$ . Then for any  $t > 0$ ,

$$P(|f(X_1, \dots, X_n) - \mathbb{E}f(X_1, \dots, X_n)| \geq t) \leq 2e^{-\frac{2t^2}{\sum_{i=1}^n c_i^2}}$$



pf: Let  $\underline{X} = (X_1, \dots, X_n)$ ,  $Z_0 = \mathbb{E} f(\underline{X})$ ,  $Z_i = \mathbb{E}(f(\underline{X}) | X_1, \dots, X_i)$ ,  $Z_n = f(\underline{X})$ .

Claim  $\mathbb{E}(e^{\lambda(Z_k - Z_{k-1})} | X_1, \dots, X_{k-1}) \leq e^{\frac{\rho^2 c_k^2}{8}}$ ,  $1 \leq k \leq n$ .

pf: Let  $U_k = \sup_u \{ \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}, u) - \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}) \}$

$L_k = \inf_l \{ \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}, l) - \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}) \}$

Note that  $U_k L_k \leq \sup_{l, u} \{ \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}, u) - \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{k-1}, l) \}$

$\stackrel{\substack{\because X_1, \dots, X_n \\ \text{are independent}}}{\leq} \sup_{l, u} \left\{ \sum_{j=k+1}^n [f(X_1, \dots, X_{k-1}, u, y_{k+1}, \dots, y_n) - f(X_1, \dots, X_{k-1}, l, y_{k+1}, \dots, y_n)] \prod_{j=k+1}^n P(X_j = y_j) \right\}$

$\leq C_k$  ( $\because$  Lipschitz condition)

So  $L_k \leq Z_k - Z_{k-1} \leq U_k \leq L_k + C$  and hence the claim is true by the extension of Hoeffding's Lemma, since  $\mathbb{E}(Z_k - Z_{k-1} | X_1, \dots, X_{k-1}) = 0$  a.s.

QED of claim

Pf (continued)

$$P\{f(\underline{x}) - \varepsilon f(\underline{x}) \geq t\}$$

$$\leq e^{-\rho t} \varepsilon e^{\rho [f(\underline{x}) - \varepsilon f(\underline{x})]}$$

$$= e^{-\rho t} \varepsilon e^{\rho \sum_{k=1}^n [z_k - z_{k-1}]}$$

$$= e^{-\rho t} \varepsilon \left( \varepsilon \left( e^{\rho \sum_{k=1}^n [z_k - z_{k-1}]} \mid X_1, \dots, X_{n-1} \right) \right) \quad (\because \text{Tower property})$$

$$= e^{-\rho t} \varepsilon \left( e^{\rho \sum_{k=1}^{n-1} [z_k - z_{k-1}]} \varepsilon \left( e^{\rho (z_n - z_{n-1})} \mid X_1, \dots, X_{n-1} \right) \right)$$

$$\leq e^{-\rho t} \varepsilon \left( e^{\rho \sum_{k=1}^{n-1} [z_k - z_{k-1}]} e^{\frac{\rho^2 C_n^2}{8}} \right) \quad (\text{by the claim})$$

$$\leq e^{-\rho t} \prod_{k=1}^n e^{\frac{\rho^2 C_k^2}{8}} \quad (\text{by repeating the same argument } n \text{ times})$$

$$= e^{-\rho t + \rho^2 \sum_{k=1}^n \frac{C_k^2}{8}} \leq e^{-\frac{2t^2}{\sum_{k=1}^n C_k^2}} \quad \text{by choosing } \rho = \frac{4t}{\sum_{k=1}^n C_k^2}$$

**QED**



# Bin Packing

- The **Bin Packing Problem** requires finding the minimum number of unit size bins needed to pack a given collection of items with sizes in  $[0,1]$ . **In our model, we have  $n$  items to pack**, and the size of items are  $X_1, \dots, X_n$  i.i.d. over  $[0,1]$ . *We use a fixed procedure for packing.*

**Ref:**

Rhee & Talagrand (1987) "Martingale inequalities and NP-complete problem"  
Math. of Oper. Res., 12, 177-181.

# Bin Packing

Discussion: We denote by  $f(x_1, \dots, x_n)$  the number of bins needed to pack  $x_1, x_2, \dots, x_n$  using procedure  $\mathcal{A}$ .

(a) If  $\mathcal{A}$  is optimum then  $f$  is 1-Lipschitz. McDiarmid's Ineq.

says that 
$$P(|f(\underline{x}) - \mathbb{E}f(\underline{x})| \geq t) \leq 2e^{-\frac{t^2}{n}}.$$

(b) If  $\mathcal{A}$  is the procedure Next Fit then  $f$  is 2-Lipschitz,

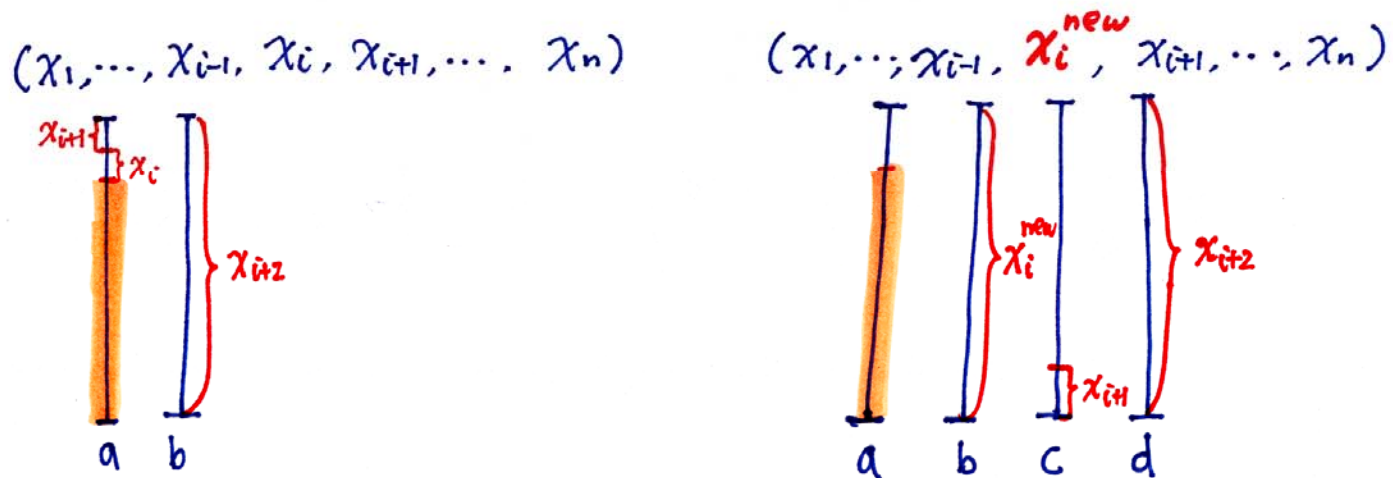
and hence 
$$P(|f(\underline{x}) - \mathbb{E}f(\underline{x})| \geq t) \leq 2e^{-\frac{t^2}{n}}.$$



# Next Fit Procedure

- Where the bins are filled one at a time and a new bin is started when the current element does not fit in the bin being currently filled.

To see  $f$  is 2-Lipschitz:



# Concentration of the Chromatic Number

Thm (Shamir-Spencer) Let  $n \geq 2$  and  $p \in (0, 1)$ .

Then we have

$$P(|\chi(G_{n,p}) - \mathbb{E}\chi(G_{n,p})| \geq t) \leq 2e^{-\frac{2t^2}{n-1}}$$

Remark: probability space  $\Omega = \{\omega : \omega \text{ is a graph on the vertex set } \{1, 2, \dots, n\}\}$

with probability measure  $P(\omega) = p^m (1-p)^{\binom{n}{2}-m}$  where  $m = |E(\omega)|$ .

One may consider  $G_{n,p}$  as a **random element**  $G_{n,p} : \Omega \rightarrow \Omega$  s.t.  $G_{n,p}(\omega) = \omega$ .  
Sometimes we consider  $G_{n,p}$  as a prob. space.



**pf:** let  $\underline{x}_1 = (x_{12}, x_{13}, x_{14}, \dots, x_{1n})$

$\underline{x}_2 = (x_{23}, x_{24}, \dots, x_{2n})$

$\vdots$

$\underline{x}_{n-2} = (x_{n-2,n-1}, x_{n-2,n})$

$\underline{x}_{n-1} = x_{n-1,n}$

We have  $\chi(G_{n,p}) = f(\underline{x}_1, \underline{x}_2, \dots, \underline{x}_{n-1})$  for some function  $f$ .

Note that  $f: S_1 \times S_2 \times \dots \times S_{n-1} \rightarrow \mathbb{R}$ , where  $S_i = \{0,1\}^{n-i}$ , and  $f$  is a 1-Lipschitz function.

By McDiarmid's Inequality, we arrive at

$$P(|\chi(G_{n,p}) - \mathbb{E}\chi(G_{n,p})| \geq t) \leq 2e^{-\frac{2t^2}{n-1}}$$

**QED**

## A Remark on Shamir-Spencer's Thm.

- Since  $P(|\chi(G_{n,p}) - \mathbb{E} \chi(G_{n,p})| \geq t\sqrt{n-1}) \leq 2e^{-2t^2}$ , so the chromatic number is almost always concentrated on about  $O(\sqrt{n})$  values.
- Note that we have no clue to what the value of  $\mathbb{E} \chi(G_{n,p})$  is.



# A technique lemma...

Thm Lemma 7.3.4 (Alon & Spencer) Let  $\alpha, c$  be fixed and  $\alpha > \frac{5}{6}$ . Let  $p = n^{-\alpha}$ . Then almost always every  $c\sqrt{n \log n}$  vertices of  $G_{n,p}$  induce a 3-colorable subgraph.

$$\text{pf: } P(\forall \text{ subgraph } H \text{ of } G_{n,p} \text{ with } \nu_H \leq c\sqrt{n \log n} \text{ having } \chi(H) \leq 3) \\ = 1 - P\left(\bigcup_{t=4}^{c\sqrt{n \log n}} \{\exists H \subseteq G_{n,p} \text{ with } \nu_H = t \text{ having } \chi(H) > 3\}\right)$$

$$\begin{aligned} & \leq \sum_{t=4}^{c\sqrt{n \log n}} P(\exists \text{ minimal } H \subseteq G_{n,p} \text{ with } \nu_H = t \text{ having } \chi(H) > 3) \\ & \quad \text{i.e. } \chi(H-v) \leq 3 \text{ for } \forall v \in V(H) \\ & \leq \sum_{t=4}^{c\sqrt{n \log n}} P(\exists H \subseteq G_{n,p} \text{ with } \nu_H = t \text{ having } d_H(x) \geq 3 \text{ for all } x \in V(H)) \\ & \leq \sum_{t=4}^{c\sqrt{n \log n}} \binom{n}{t} P(H \cong G_{n,p}[\{1, 2, \dots, t\}] \text{ has } e_H \geq \frac{3}{2}t) \end{aligned}$$

pf (continued)

$$\leq \sum_{t=4}^{c\sqrt{n \log n}} \binom{n}{t} \binom{\binom{t}{2}}{\frac{3}{2}t} p^{\frac{3}{2}t}$$

$$\leq \sum_{t=4}^{c\sqrt{n \log n}} \left( \frac{en}{t} \right)^t \left( \frac{e \binom{t}{2}}{\frac{3}{2}t} \right)^{\frac{3}{2}t} p^{\frac{3}{2}t} \quad (\because \binom{n}{k} \leq \left( \frac{en}{k} \right)^k)$$

$$\leq \sum_{t=4}^{c\sqrt{n \log n}} \left( \frac{en}{t} \frac{t^{\frac{3}{2}} e^{\frac{3}{2}}}{3^{\frac{3}{2}}} n^{-\frac{3}{2}\alpha} \right)^t \quad (\because p = n^{-\alpha})$$

$$\leq \sum_{t=4}^{c\sqrt{n \log n}} \left( c_1 n^{1-\frac{3}{2}\alpha} t^{\frac{1}{2}} \right)^t$$

$$\leq \sum_{t=4}^{c\sqrt{n \log n}} \left( c_2 n^{1-\frac{3}{2}\alpha} n^{\frac{1}{4}} (\log n)^{\frac{1}{4}} \right)^t$$

$$= \sum_{t=4}^{c\sqrt{n \log n}} \left( c_2 n^{-\epsilon} (\log n)^{\frac{1}{4}} \right)^t \quad (\because \frac{5}{4} - \frac{3}{2}\alpha < 0 \iff \frac{5}{6} < \alpha)$$

$$\leq \left( c_2 n^{-\epsilon} (\log n)^{\frac{1}{4}} \right)^4 \frac{1}{1 - c_2 n^{-\epsilon} (\log n)^{\frac{1}{4}}} = o(1)$$

**QED**



# Four-Value Concentration

Thm (see Alon & Spencer, Thm 7.3.3)

Let  $\alpha > \frac{5}{6}$  be fixed, and let  $p = n^{-\alpha}$  (i.e.  $p$  is not too large)

Then for any  $n$ ,  $\exists u = u(\alpha, n)$  such that

$\chi(G_{n,p}) \in \{u, u+1, u+2, u+3\}$  almost surely

i.e.  $\Pr\{\chi(G_{n,p}) \notin \{u, u+1, u+2, u+3\}\} \rightarrow 0$  as  $n \rightarrow \infty$

pf: let  $u = u(n, \alpha)$  be the smallest integer st.  $P(X(G_{n,p}) \leq u) > \frac{1}{n}$ .

Claim A:  $P(X(G_{n,p}) \geq u) \geq 1 - \frac{1}{n}$

pf: The choice of  $u \Rightarrow P(X(G_{n,p}) \leq u-1) \leq \frac{1}{n} \Rightarrow P(X(G_{n,p}) \geq u) \geq 1 - \frac{1}{n}$  **IA**

Let  $X$  be the minimum number of vertices whose deletion makes  $G_{n,p}$   $u$ -colorable.

Then  $X = f(\underline{x}_1, \underline{x}_2, \dots, \underline{x}_{n-1})$  for some function  $f$ , where  $\underline{x}_1, \underline{x}_2, \dots, \underline{x}_{n-1}$  were defined in the proof of Shamir-Spencer's Thm. Note that  $f$  is  $1$ -Lipschitz.

Claim B:  $\sqrt{2(n-1) \log n} > \epsilon X$

$$\begin{aligned} \text{pf: } \frac{1}{n} &< P(X(G_{n,p}) \leq u) = P(X=0) \\ &= P(X \leq \epsilon X - \epsilon X) \\ &= P(f(\underline{x}_1, \dots, \underline{x}_{n-1}) \leq \epsilon f(\underline{x}_1, \dots, \underline{x}_{n-1}) - \epsilon X) \\ &\leq \exp\left(-\frac{2(\epsilon X)^2}{n-1}\right) \text{ by McDiarmid's Inequality} \end{aligned}$$

Therefore  $\sqrt{\frac{1}{2}(n-1) \log n} > \epsilon X$ .

**IB**



## pf (Continued)

Claim C:  $P(X < 2\sqrt{2(n-1)\log n}) \geq 1 - \frac{1}{n}$

pf: LHS =  $1 - P(X \geq 2\sqrt{2(n-1)\log n})$

$$\geq 1 - P(X \geq EX + \sqrt{2(n-1)\log n})$$

$$\geq 1 - \exp\left(-\frac{2(n-1)\log n}{2(n-1)}\right) \quad \text{by McDiarmid's Ineq. again!}$$

$$= 1 - \frac{1}{n}. \quad \text{IC}$$

Let  $A = \{X(G_{n,p}) \geq u\}$  and  $B = \{X < 2\sqrt{2(n-1)\log n}\}$  be two events.

$$\begin{aligned} \text{Then } P(A \cap B) &= 1 - P(\bar{A} \cup \bar{B}) \geq 1 - P(\bar{A}) - P(\bar{B}) \\ &\geq 1 - \frac{1}{n} - \frac{1}{n} = 1 - \frac{2}{n} \rightarrow 1 \text{ as } n \rightarrow \infty \end{aligned}$$

The thm. follows from the technique lemma we proved.

**QED**

# Kim's Lemma

Thm (J.H. Kim 1995) Suppose that

- $X_1, X_2, \dots, X_n$  are independent rvs s.t.  $X_i \sim B(1, p_i) \forall i$ .
- $f: \{0,1\}^n \rightarrow \mathbb{R}$  and  $g$  is a convex function.

Then for any  $i$ ,  $1 \leq i \leq n$ , we have

$$\mathbb{E}(g(V_i) | X_1, \dots, X_{i-1}) \leq \mathbb{E}\left(p_i g(p_i \gamma_i) + \overset{q_i = 1-p_i}{q_i} g(-p_i \gamma_i) \mid X_1, \dots, X_{i-1}\right)$$

Where  $\gamma_i = f(X_1, \dots, X_{i-1}, 1, X_{i+1}, \dots, X_n) - f(X_1, \dots, X_{i-1}, 0, X_{i+1}, \dots, X_n)$

$$V_i = \mathbb{E}(f(\underline{X}) | X_1, \dots, X_i) - \mathbb{E}(f(\underline{X}) | X_1, \dots, X_{i-1})$$

$\underline{X} = (X_1, \dots, X_n)$



# An extension of Devroye's Inequality

Thm If we have

(1)  $X_1, X_2, \dots, X_n$  are independent,  $X_i \sim B(1, p_i) \quad \forall i$

(2)  $f: \{0, 1\}^n \rightarrow \mathbb{R}$  having a vector  $(c_1, c_2, \dots, c_n)$  s.t.

$|f(\underline{a}) - f(\underline{b})| \leq c_i$  for all  $\underline{a}, \underline{b}$  that differ only in the  $i$ th coordinate.

Then

$$\text{Var}[f(X_1, X_2, \dots, X_n)] \leq \sum_{i=1}^n c_i^2 p_i (1 - p_i)$$

pf: Let  $f = f(X_1, X_2, \dots, X_n)$  and  $V_i = \mathbb{E}(f | \underbrace{X_1, \dots, X_i}_{X_i}) - \mathbb{E}(f | \underbrace{X_1, \dots, X_{i-1}}_{X_{i-1}})$   $2 \leq i \leq n$

Then  $\text{Var } f = \mathbb{E}(f - \mathbb{E}f)^2$

$V_1 \stackrel{\text{def}}{=} \mathbb{E}(f | X_1) - \mathbb{E}(f | \mathcal{F}_0)$

$$= \mathbb{E}\left(\left(\sum_{i=1}^n V_i\right)^2\right)$$

$\because \mathbb{E}(f | \mathcal{F}_0) = \mathbb{E}f$  where  $\mathcal{F}_0 = \{\emptyset, \Omega\}$   
 $\mathbb{E}(f | X_n) = f$

$$= \mathbb{E}\left(\sum_{i=1}^n V_i^2 + 2 \sum_{1 \leq i < j \leq n} V_i V_j\right)$$

$$= \mathbb{E}\left(\sum_{i=1}^n V_i^2\right)$$

$\because \mathbb{E} V_i V_j = \mathbb{E}(\mathbb{E}(V_i V_j | X_1, \dots, X_{j-1}))$   $1 \leq i < j \leq n$   
 $= \mathbb{E}(V_i \mathbb{E}(V_j | X_1, \dots, X_{j-1})) = \mathbb{E}(V_i \cdot 0) = 0$

$$= \sum_{i=1}^n \mathbb{E}(\mathbb{E}(V_i^2 | \underline{X}_{i-1}))$$

here we define  $\mathbb{E}(V_i^2 | \underline{X}_0) \stackrel{\text{def}}{=} \mathbb{E}(V_i^2 | \mathcal{F}_0)$

$$\leq \sum_{i=1}^n \mathbb{E}\left(\mathbb{E}(P_i (g_i(r_i))^2 + g_i(P_i r_i)^2 | \underline{X}_{i-1})\right)$$

Applying Kim's lemma  
 with  $g(x) = x^2$ .

$$= \sum_{i=1}^n P_i g_i r_i^2$$

$$\leq \sum_{i=1}^n P_i g_i C_i^2$$

$\because |r_i| \leq C_i$  Lipschitz property

**QED**



- 1
- $k$ - $L(2,1)$ -<sup>coloring</sup>labeling of  $G$ :

$\varphi: V(G) \rightarrow \{0, 1, 2, \dots, k\}$  such that

$$x \sim y \Rightarrow |\varphi(x) - \varphi(y)| \geq 2$$

$$x \wedge y \Rightarrow |\varphi(x) - \varphi(y)| \geq 1$$

- $\lambda_{2,1}(G) \stackrel{\text{def}}{=} \min \{k: G \text{ has a } k\text{-}L(2,1)\text{-}\sup_{\text{coloring}}\text{labeling}\}$

- $k$ - $L(a,b)$ -<sup>coloring</sup>labeling of  $G$ :

$\varphi: V(G) \longrightarrow \{0, 1, 2, \dots, k\}$  such that

$$x \sim y \Rightarrow |\varphi(x) - \varphi(y)| \geq a$$

$$x \triangle y \Rightarrow |\varphi(x) - \varphi(y)| \geq b$$

- $\lambda_{a,b}(G) \stackrel{\text{def}}{=} \min \{k : G \text{ has a } k\text{-}L(a,b)\text{-}\sup_{\text{coloring}}\text{labeling}\}$



# Something New

Theorem Let  $n \geq 2$  and  $p \in (0,1)$  be arbitrary,  
and let  $c = E[\lambda_{2,1}(G_{n,p})]$ .

Then

$$\Pr\left\{|\lambda_{2,1}(G_{n,p}) - c| \geq t\sqrt{n(n-1)}\right\} \leq 2e^{-t^2}$$

Pf:

- $\lambda_{2,1}(G_{n,p})$  can be expressed as a function of  $X_{e_1}, \dots, X_{e_m}$ ,  $m = \binom{n}{2}$
- $\lambda_{2,1}(G_{n,p})$  is a 2-Lipschitz function in the  $X_{e_i}$ 's
- McDiarmid's inequality now gives what we want.

# Another Something New

Thm For any  $n$ , there is a  $u=u(n)$   
such that

$$\lambda_{2,1}(G_{n,p}) \in [u, u + 4\sqrt{n(n-1)\log n} + 1]$$

almost surely.



# Sparse Random Graphs

Observation: If  $p \leq 1/n \log n$  and  $c = E[\lambda_{2,1}(G_{n,p})]$   
then

$$\begin{aligned} & \Pr\{|\lambda_{2,1}(G_{n,p}) - c| \geq \sqrt{n}\} \\ & \leq \frac{\text{Var}\{\lambda_{2,1}(G_{n,p})\}}{n} \\ & \leq \frac{\sum_{i=1}^{\binom{n}{2}} 4p(1-p)}{n} \\ & \leq \frac{2}{\log n} \rightarrow 0 \text{ as } n \rightarrow \infty \end{aligned}$$



Thank you